

Trade and Inequality: From Theory to Estimation

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Motivation

- Neoclassical trade theory (H-O, SF, R)
 - sector-level comparative advantage
 - focus on “**between**” effects
- New trade theory
 - Krugman: intra-industry trade
 - Melitz: firm-level comparative advantage
 - focus on “**within**” effects
- Trade and inequality
 - Heavily influenced by H-O framework
 - Empirically has limited explanatory power
- “New view” of trade and inequality
 - link wages to firm performance
 - **within**-industry, **between**-firm

This Paper

- ① Uses matched employer-employee data from Brazil for 1986-1998
 - increase in inequality [▶ show](#)
 - trade liberalization [▶ show](#)
- ② Documents, for both levels and changes, the importance of inequality:
 - within sector-occupations
 - for workers with similar observables (residual inequality)
 - between firms
- ③ Structurally estimates a heterogenous-firm model of firm employment, wages and export status
 - extension of HIR (2010)
 - a model of *within-sector*, *between-firm residual* inequality
 - reproduces empirical size and exporter wage premia:

$$w_{ij} = \alpha + \beta h_j + \delta_{xj} + \dots + \varepsilon_{ij}$$

Related Literature

- Long and large tradition in labor literature
- “New view” empirics:
 - Bernard and Jensen (1995)...
 - Verhoogen (2008)
 - Amity and Davis (2011)
 - AKM (1999) estimation used in trade context
- “New view” theory:
 - Feenstra and Hanson (1999)...
 - Yeaple (2005)...
 - Egger and Kreickemeier (2009)
 - HIR (2010)...

DATA

Brazilian RAIS Data

- Matched employer-employee data from 1986–1998
 - All workers employed in the formal sector
 - Focus on the manufacturing sector
 - Observe firm, industry and occupation
 - Observe worker education (high school, college degree), demographics (age, sex) and experience (employment history)
- Over the period 1986-1998 as a whole, our sample includes more than 7 million workers and 100,000 firms in every year
- Trade transactions data from 1986-1998
 - Merged with the matched employer-employee data
 - Observe firm exports and export products and destinations

Sectors

- Twelve aggregate sectors (IBGE) 1986-1998

| | Industry | Empl't share (%) | Rel. mean log wage | Fraction Exporters | Exporter Empl't % |
|----|--------------------------------|-------------------------|---------------------------|---------------------------|--------------------------|
| 2 | Non-metallic mineral products | 5.5 | -0.12 | 2.3 | 32.3 |
| 3 | Metallic products | 9.8 | 0.27 | 6.1 | 49.9 |
| 4 | Machinery, equipment & instr. | 6.6 | 0.38 | 12.3 | 54.1 |
| 5 | Electrical & telecomm. equip. | 6.0 | 0.37 | 11.8 | 56.3 |
| 6 | Transport equipment | 6.3 | 0.61 | 11.2 | 70.6 |
| 7 | Wood products & furniture | 6.5 | -0.48 | 3.2 | 23.5 |
| 8 | Paper, publishing & printing | 5.4 | 0.14 | 2.5 | 30.6 |
| 9 | Rubber, tobacco, leather & fur | 7.0 | -0.04 | 8.6 | 50.8 |
| 10 | Chemical & pharma. products | 9.9 | 0.40 | 11.2 | 50.6 |
| 11 | Apparel & textiles | 15.7 | -0.32 | 2.5 | 34.8 |
| 12 | Footwear | 4.4 | -0.44 | 12.2 | 65.7 |
| 13 | Food, beverages & alcohol | 16.9 | -0.30 | 3.9 | 38.0 |

- More than 250 disaggregated industries (CNAE) 1994-1998

Occupations

- Five aggregate occupations 1986-1998

| | Occupation | Employment share (%) | Relative mean log wage |
|---|-----------------------------|---------------------------------|-----------------------------------|
| 1 | Professional and Managerial | 7.8 | 1.08 |
| 2 | Skilled White Collar | 11.1 | 0.40 |
| 3 | Unskilled White Collar | 8.4 | 0.13 |
| 4 | Skilled Blue Collar | 57.4 | -0.15 |
| 5 | Unskilled Blue Collar | 15.2 | -0.35 |

- More than 300 disaggregated occupations (CBO) 1986-1998

STYLIZED FACTS

Within and Between Inequality

Sector-occupation bins

| | Level (%) | Change (%) |
|--|-----------|------------|
| A. Main Period | 1994 | 1986–95 |
| Within occupation | 82 | 92 |
| Within sector | 83 | 73 |
| Within sector-occupation | 68 | 66 |
| Within detailed-occupation | 61 | 60 |
| Within sector–detailed-occupation | 56 | 54 |
| B. Late Period | 1994 | 1994–98 |
| Within detailed-sector –detailed-occupation | 47 | 141 |

Fact 1

Within sector-occupation component of wage inequality accounts for over 2/3 of both level and growth of wage inequality

Residual Inequality

Conditional on worker observables

| | Level (%) 1994 | Change (%) 1986–95 |
|----------------------------|-------------------|-----------------------|
| Residual wage inequality | 59 | 49 |
| — within sector-occupation | 89 | 91 |

Fact 2

- (i) *Residual inequality is at least as important as worker observables for both level and growth of wage inequality*
- (ii) *Almost all residual inequality is within sector-occupations*

Between-firm Inequality

- Mincer log-wage regression with firm fixed effect:

$$w_{it} = z'_{it} \vartheta_{\ell t} + \psi_{j\ell t} + \nu_{it}$$

- i worker
- j firm
- ℓ sector-occupation bin
- $\psi_{j\ell t}$ firm fixed effect includes:
 - Returns to unobserved skill (workforce composition)
 - Worker rents (differences in wage for same workers)
 - Match effects
- Decomposition of within inequality:
 - Observables: $\text{var}(z'_{it} \hat{\vartheta}_{\ell t})$
 - Between-firm component: $\text{var}(\hat{\psi}_{j\ell t})$
 - Covariance: $\text{cov}(z'_{it} \hat{\vartheta}_{\ell t}, \hat{\psi}_{j\ell t})$
 - Within-firm component: $\text{var}(\hat{\nu}_{it})$

Between-firm Inequality

Within sector-occupation bins

| | UNCONDITIONAL FIRM WAGE COMPONENT, $\psi_{j\ell t}^U$ | | CONDITIONAL FIRM WAGE COMPONENT, $\psi_{j\ell t}^C$ | |
|--------------------------------|---|-------------------------|---|-------------------------|
| | Level (%) 1994 | Change (%) 1986–1995 | Level (%) 1994 | Change (%) 1986–1995 |
| Between-firm wage inequality | 55 | 115 | 39 | 86 |
| Within-firm wage inequality | 45 | -15 | 37 | -11 |
| Worker observables | | | 13 | 2 |
| Covar observables–firm effects | | | 11 | 24 |

Fact 3

Between-firm component account for about half of level and the majority of growth of within sector-occupation wage inequality

Between-firm Inequality

Size and exporter wage premia

| | UNCONDITIONAL FIRM WAGE COMPONENT, $\hat{\psi}_{jt}^U$ | CONDITIONAL FIRM WAGE COMPONENT, $\hat{\psi}_{jt}^C$ |
|----------------------|--|--|
| Firm Employment Size | 0.122*** (0.010) | 0.104*** (0.009) |
| Firm Export Status | 0.262*** (0.042) | 0.168*** (0.024) |
| Sector Fixed Effects | yes | yes |
| Within R-squared | 0.17 | 0.13 |
| Observations | 91, 410 | 91, 410 |

Fact 4

Larger firms on average pay higher wages; exporters on average pay higher wages even after controlling for size. The remaining variation in wages is substantial.

STRUCTURAL MODEL

Model: Extension of HIR

- ① Melitz (2003) product market:

$$R = \Upsilon A y^\beta, \quad \Upsilon \in \{1, \Upsilon_x > 1\}$$

- ② Heterogeneity in fixed cost of exports: $e^\varepsilon F_x$

- ③ Complementarity between productivity and worker ability:

$$y = e^\theta H^\gamma \bar{a}, \quad \gamma < 1$$

- ④ Unobserved heterogeneity and costly screening:

$$e^{-\eta} C \frac{(a_c)^\delta}{\delta} \Rightarrow \bar{a} = \frac{k}{k-1} a_c$$

- ⑤ DMP frictional labor market and wage bargaining:

$$W = \frac{\beta\gamma}{1 + \beta\gamma} \frac{R}{H} = b \cdot (a_c)^{k/\delta}$$

Model Predictions

- A firm with idiosyncratic shock $\{\theta, \eta, \varepsilon\}$:

$$R(\theta, \eta, \varepsilon) = \kappa_r \Upsilon^{\frac{1-\beta}{\Gamma}} (e^\theta)^{\frac{\beta}{\Gamma}} (e^\eta)^{\frac{\beta(1-\gamma k)}{\delta \Gamma}}$$

$$H(\theta, \eta, \varepsilon) = \kappa_h \Upsilon^{\frac{(1-\beta)(1-k/\delta)}{\Gamma}} (e^\theta)^{\frac{\beta(1-k/\delta)}{\Gamma}} (e^\eta)^{\frac{\beta(1-\gamma k)(1-k/\delta)}{\delta \Gamma} - \frac{k}{\delta}}$$

$$W(\theta, \eta, \varepsilon) = \kappa_w \Upsilon^{\frac{k(1-\beta)}{\delta \Gamma}} (e^\theta)^{\frac{\beta k}{\delta \Gamma}} (e^\eta)^{\frac{k}{\delta} \left(1 + \frac{\beta(1-\gamma k)}{\delta \Gamma}\right)}$$

- Market access variable

$$\Upsilon = 1 + \iota \cdot (\Upsilon_x - 1), \quad \Upsilon_x = 1 + \tau^{-\frac{\beta}{1-\beta}} \frac{A_x}{A_d}$$

- Selection into exporting

$$\iota = \iota(\theta, \eta, \varepsilon) = \mathbb{I} \left\{ \kappa_\pi \left(\Upsilon_x^{\frac{1-\beta}{\Gamma}} - 1 \right) (e^\theta)^{\frac{\beta}{\Gamma}} (e^\eta)^{\frac{\beta(1-\gamma k)}{\delta \Gamma}} \geq F_x e^\varepsilon \right\}$$

Econometric Model

- Empirical model of $X_j = \{h_j, w_j, \iota_j\}_j$:

$$\begin{cases} h_j = \alpha_h + \mu_h \cdot \iota_j + u_j, \\ w_j = \alpha_w + \mu_w \cdot \iota_j + \zeta u_j + v_j, \\ \iota_j = \mathbb{I}\{z_j \geq f\} \end{cases}$$

- Distributional assumption:

$$(u_j, v_j, z_j)' \sim \mathcal{N}(\mathbf{0}, \Sigma), \quad \Sigma = \begin{pmatrix} \sigma_u^2 & & & \\ 0 & \sigma_v^2 & & \\ \rho_u \cdot \sigma_u & \rho_v \cdot \sigma_v & & 1 \end{pmatrix}$$

- **Selection** (ρ_u, ρ_v) versus **Market access** (μ_h, μ_w)
- Theoretical restriction: $\mu_h, \mu_w > 0$

Identification

① Maximum Likelihood

- under additional orthogonality assumption between structural productivity shocks θ and η :

$$\zeta \leq \frac{\mu_w}{\mu_h} \leq \zeta + \frac{\sigma_v^2}{(1 + \zeta)\sigma_u^2}$$

② GMM Bounds

- based on a subset of moments

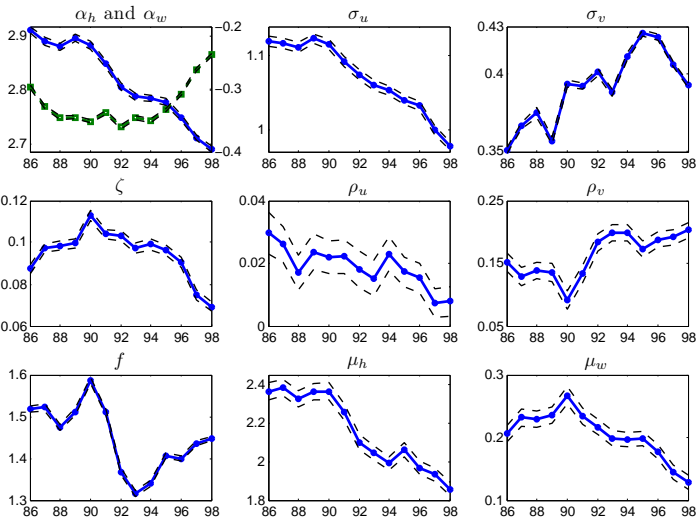
③ Semi-parametric estimation

- using alternative instruments for export participation

RESULTS

Estimation Results

Parameters

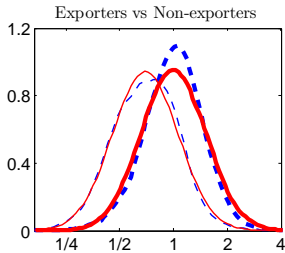
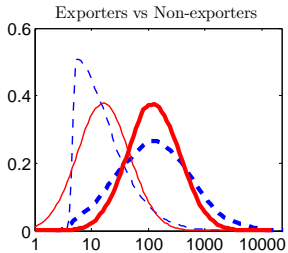
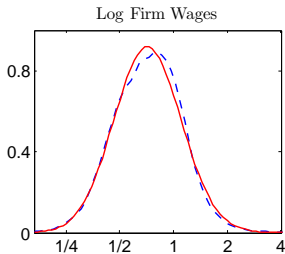
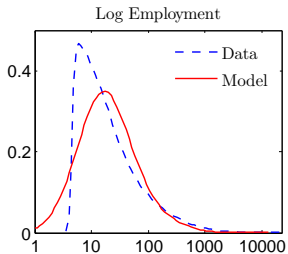


Model Fit

Firm-level moments

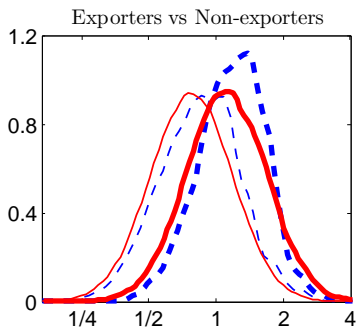
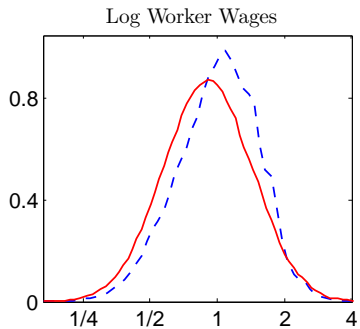
| | All firms | Non-exp. | Exporters |
|-----------------------|-----------|----------|-----------|
| | DATA | | |
| Mean h | 2.96 | 2.78 | 4.82 |
| Mean w | -0.33 | -0.37 | -0.01 |
| Std deviation h | 1.20 | 1.00 | 1.46 |
| Std deviation w | 0.43 | 0.43 | 0.38 |
| Correlation h & w | 0.33 | 0.24 | 0.32 |
| Fraction of exporters | 9.0% | | |
| | MODEL | | |
| Mean h | 2.96 | 2.78 | 4.83 |
| Mean w | -0.33 | -0.37 | 0.00 |
| Std deviation h | 1.20 | 1.05 | 1.05 |
| Std deviation w | 0.43 | 0.42 | 0.42 |
| Correlation h & w | 0.32 | 0.25 | 0.24 |
| Fraction of exporters | 9.0% | | |

Employment and Wage Distributions



--- Data (non-exp) --- Data (exporters) — Model (non-exp) — Model (exporters)

Worker Wage Distribution



Model Fit

Worker wage dispersion

| | DATA | MODEL |
|------------------|------|-------|
| Std deviation | 0.42 | 0.46 |
| — non-exporters | 0.42 | 0.42 |
| — exporters | 0.35 | 0.42 |
| Gini coefficient | 0.23 | 0.25 |
| 90/10-ratio | 2.95 | 3.23 |
| — 90/50 | 1.63 | 1.80 |
| — 50/10 | 1.81 | 1.80 |

Model Fit

Worker wage dispersion

| | DATA | MODEL |
|------------------|------|-------|
| Std deviation | 0.42 | 0.46 |
| — non-exporters | 0.42 | 0.42 |
| — exporters | 0.35 | 0.42 |
| Gini coefficient | 0.23 | 0.25 |
| 90/10-ratio | 2.95 | 3.23 |
| — 90/50 | 1.63 | 1.80 |
| — 50/10 | 1.81 | 1.80 |

Size and exporter wage premia

| | DATA | MODEL |
|--------------------|------|-------|
| Employment premium | 0.10 | 0.10 |
| Exporter premium | 0.16 | 0.16 |
| <i>R</i> -squared | 0.11 | 0.11 |

Counterfactuals

- Estimated model:

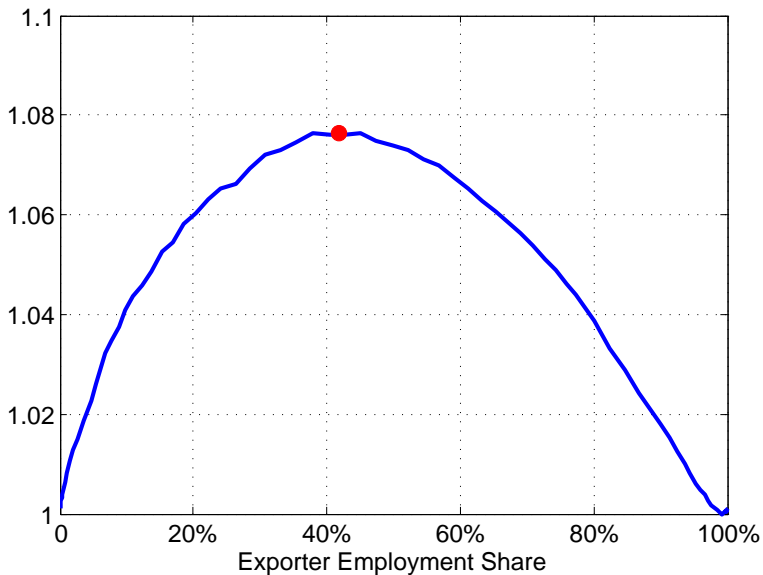
$$\begin{cases} h_j = \alpha_h + \mu_h \cdot l_j + u_j, \\ w_j = \alpha_w + \mu_w \cdot l_j + \zeta u_j + v_j, \\ l_j = \mathbb{I}\{z_j \geq f\} \end{cases} \quad (u_j, v_j, z_j)' \sim \mathcal{N}(\mathbf{0}, \Sigma)$$

- Parameters (μ_h, μ_w, f) form a **sufficient statistic**:

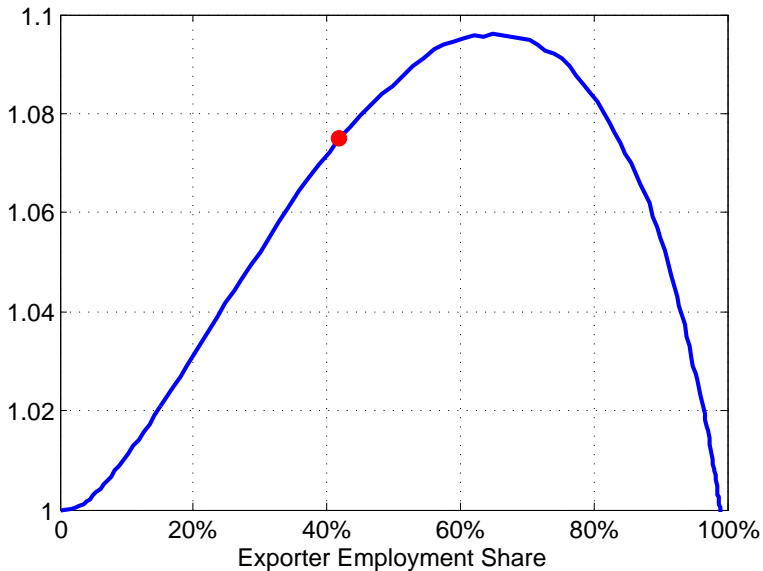
$$f = \frac{1}{\sigma} \left[\alpha_f + \log F_x - \log \left(\Upsilon_x^{\frac{1-\beta}{\Gamma}} - 1 \right) \right]$$
$$\mu_h + \mu_w = \Upsilon_x^{\frac{1-\beta}{\Gamma}}, \quad \Upsilon_x = 1 + \tau^{-\frac{\beta}{1-\beta}} \frac{A_x}{A_d}$$

- Two counterfactuals: variation in F_x and τ

Variation in Fixed Export Cost



Variation in Variable Trade Cost



GMM BOUNDS

GMM Bounds

- We drop the orthogonality assumption and use the following set of moments:
 - (a) conditional first moments: $\mathbb{E}\iota$, $\mathbb{E}\{h|\iota\}$ and $\mathbb{E}\{w|\iota\}$
 - (b) unconditional second moments: $\text{var}(h)$, $\text{var}(w)$ and $\text{cov}(h, w)$
 - (c) size and exporter wage premia λ_s and λ_x , and R^2 from:

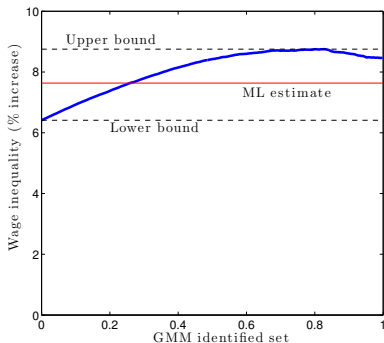
$$\mathbb{E}\{w|h, \iota\} = \lambda_0 + \lambda_s h + \lambda_x \iota$$

- In addition, we impose $|\rho_u|, |\rho_v| < 1$ and $\sigma_u, \sigma_v > 0$
- and $\text{corr}((1 + \zeta)u + v, z) = (1 + \zeta)\rho_u\sigma_u + \rho_v\sigma_v > 0$
- We check that $\mu_h, \mu_w > 0$
- This identifies a uni-dimensional interval in the 10-dimensional parameter space, the **GMM identified set** [▶ Show the iSet](#)
- For each element of this set we conduct:
 - (a) autarky and
 - (b) variable trade cost counterfactual:
 - $\tau \uparrow$ to generate a 10p.p. \downarrow in exporter employment share

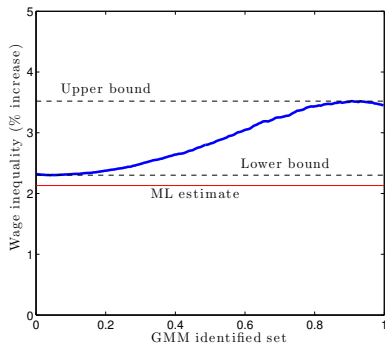
GMM Bounds

Results

(a) Autarky counterfactual



(b) τ counterfactual



- Autarky bounds: [6.6%, 9.0%] vs ML estimate 7.6%
- τ bounds: [2.3%, 3.5%] vs ML estimate 2.2%

Semi-parametric

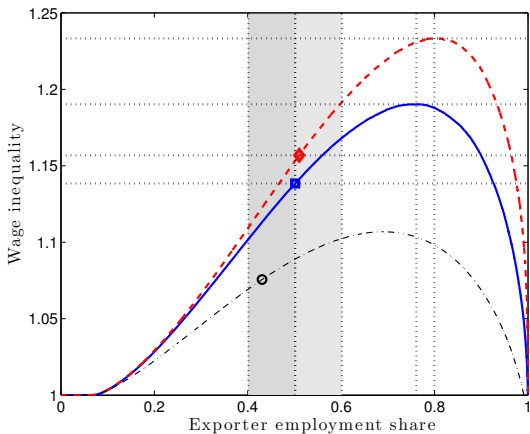
Two-stage estimation

| | (1) | (2) | (3) | (4) | (5) |
|--------------------------------------|---------------------|--------------------|---------------------------|-------------------------------------|---|
| | EASE BUSINESS | FOREIGN WORKER | MESO FOREIGN WORKER | LAYOFF MESO FOREIGN WORKER | EASE BUSINESS AND LAYOFF MESO FOREIGN WORKER |
| PANEL A: SELECTION EQUATION | | | | | |
| Ease of Business | -0.0044 (0.0009) | - | - | - | -0.0045 (0.0009) |
| Foreign Worker | - | 5.237 (0.4239) | - | - | - |
| Meso Foreign Worker | - | - | 244.9714 (10.5173) | - | - |
| Layoff Meso Foreign Worker | - | - | - | 168.2080 (10.6992) | 169.0699 |
| First-stage F-statistic | 23.88 | 152.59 | 542.53 | 247.17 | 271.67 |
| PANEL B: EMPLOYMENT EQUATION | | | | | |
| Employment market access (μ_h) | 2.0227 (0.0166) | 1.9968 (0.0168) | 2.0321 (0.0165) | 2.0324 (0.0165) | 2.0341 (0.0165) |
| Second-stage F-statistic | 52.10 | 103.65 | 23.80 | 30.26 | 14.93 |
| PANEL C: WAGE EQUATION | | | | | |
| Wage market access (μ_w) | 0.3519 (0.0045) | 0.3432 (0.0045) | 0.3121 (0.0042) | 0.3209 (0.0043) | 0.3231 (0.0044) |
| Second-stage F-statistic | 341.27 | 167.39 | 6757.53 | 4997.21 | 3964.15 |

MULTIDESTINATION

Multidestination Model

Counterfactuals

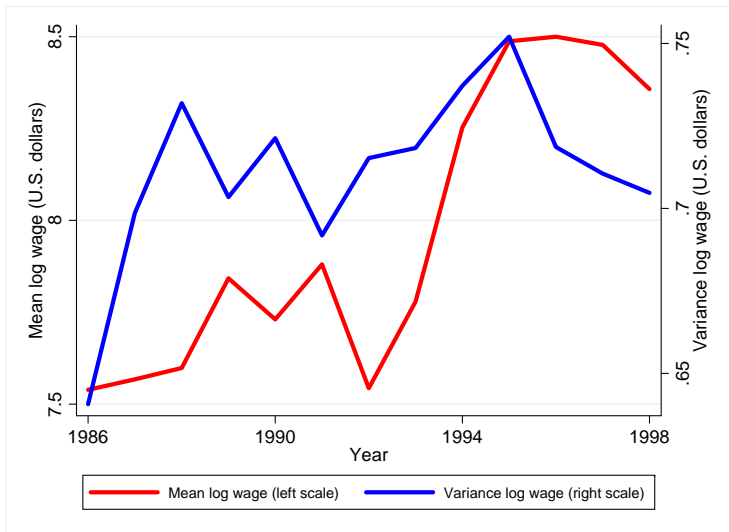


► Show the model

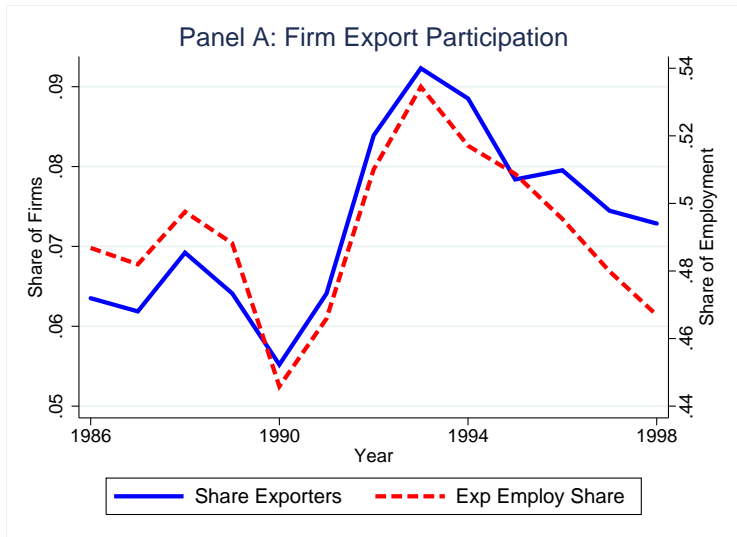
Conclusions

- Neoclassical trade theory emphasizes wage inequality between occupations and industries
- In contrast, new theories of firm heterogeneity and trade point to wage dispersion within occupations and industries
- Using matched employer-employee data for Brazil, we show:
 - Much of the increase in wage inequality since the mid-1980s has occurred **within sector-occupations**
 - Increased **within-group wage inequality**
 - Increased wage dispersion **between firms**
 - Between-firm wage dispersion related to **trade participation**
- Develop a framework for the structural estimation of a model with firm heterogeneity and wage dispersion across firms
- Use this framework to quantify the effect of trade on wage dispersion

Wage Inequality



Trade Openness



Regional Robustness

| | OVERALL | | RESIDUAL | |
|-------------------------------------|---------------|-------------------|---------------|-------------------|
| | INEQUALITY | | INEQUALITY | |
| | Level 1994 | Change 1986–95 | Level 1994 | Change 1986–95 |
| Within sector-occupation | 68 | 66 | 89 | 91 |
| Within sector-occupation, São Paulo | 64 | 49 | 89 | 71 |
| Within sector-occupation-state | 58 | 38 | 76 | 56 |
| Within sector-occupation-meso | 54 | 30 | 72 | 49 |

◀ Back to slides

Likelihood Function

$$\mathcal{L}(\Theta|X_j) = \prod_j \mathbb{P}\{(h_j, w_j, \iota_j)|\Theta\}$$

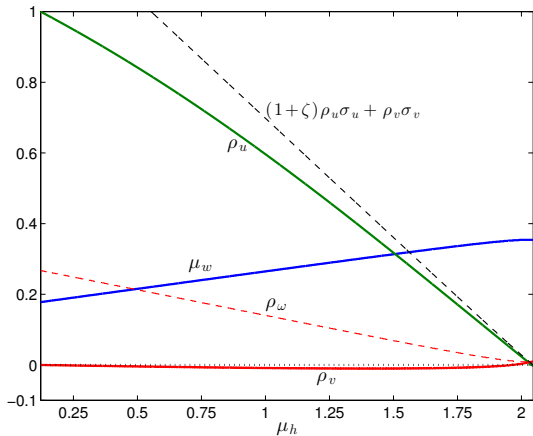
$$\mathbb{P}\{(h_j, w_j, \iota_j)|\Theta\} = \frac{1}{\sigma_u} \phi(\hat{u}_j) \frac{1}{\sigma_v} \phi(\hat{v}_j) \left[\Phi \left(\frac{f - \rho_u \hat{u}_j - \rho_v \hat{v}_j}{\sqrt{1 - \rho_u^2 - \rho_v^2}} \right) \right]^{1-\iota_j} \left[1 - \Phi \left(\frac{f - \rho_u \hat{u}_j - \rho_v \hat{v}_j}{\sqrt{1 - \rho_u^2 - \rho_v^2}} \right) \right]^{\iota_j}$$

$$\hat{u}_j = \frac{h_j - \alpha_h - \mu_h \iota_j}{\sigma_u},$$

$$\hat{v}_j = \frac{(w_j - \alpha_w - \mu_w \iota_j) - \zeta(h_j - \alpha_h - \mu_h \iota_j)}{\sigma_v}$$

GMM Identified Set

- Main idea: $\bar{h}_1 - \bar{h}_0 = \mu_h + \rho_u \sigma_u (\lambda_1 - \lambda_0)$



Multidestination Model

$$\begin{cases} h &= \alpha_h + \mu_{h,1}l_1 + (\mu_{h,2} - \mu_{h,1})l_2 + (\mu_{h,3} - \mu_{h,2})l_3 + u, \\ w &= \alpha_w + \mu_{w,1}l_1 + (\mu_{w,2} - \mu_{w,1})l_2 + (\mu_{w,3} - \mu_{w,2})l_3 + \zeta u + v, \\ l_\ell &= \mathbb{I}\{f_{\ell-1} \leq z \leq f_\ell\}, \quad \ell = 1, 2, 3, \end{cases}$$

$$\mu_{h,\ell} = \frac{\delta - k}{\delta} \log \Upsilon_{x,\ell}^{\frac{1-\beta}{\Gamma}}, \quad \mu_{w,\ell} = \frac{k}{\delta - k} \mu_{h,\ell},$$

$$f_\ell = \frac{1}{\sigma} \left[-\alpha_\pi + \log F_{x,\ell} - \log \left(\Upsilon_{x,\ell}^{\frac{1-\beta}{\Gamma}} - \Upsilon_{x,\ell-1}^{\frac{1-\beta}{\Gamma}} \right) \right],$$

$$\Upsilon_x = 1 + \tau^{-\frac{\beta}{1-\beta}} \sum_{\ell=1,2,3} l_\ell \left(\frac{A_{x,\ell}}{A_d} \right)^{\frac{1}{1-\beta}}$$

$$l_\ell = \mathbb{I} \left\{ \kappa_\pi \left[\Upsilon_{x,\ell}^{\frac{1-\beta}{\Gamma}} - \Upsilon_{x,\ell-1}^{\frac{1-\beta}{\Gamma}} \right] (e^\theta)^{\frac{\beta}{\Gamma}} (e^\eta)^{\frac{\beta(1-\gamma k)}{\delta\Gamma}} \geq e^\varepsilon F_{x,\ell} \right\}, \quad \ell = 1, 2, 3$$